Abstract—Intelligent tutoring systems adapt the curriculum to the needs of the individual student. Therefore, an accurate representation and prediction of student knowledge is essential. Bayesian Knowledge Tracing (BKT) is a popular approach for student modeling. The structure of BKT models, however, makes it impossible to represent the hierarchy and relationships between the different skills of a learning domain. Dynamic Bayesian networks (DBN) on the other hand are able to represent multiple skills jointly within one model. In this work, we suggest the use of DBNs for student modeling. We introduce a constrained optimization algorithm for parameter learning of such models. We extensively evaluate and interpret the prediction accuracy of our approach on five large-scale data sets of different learning domains such as mathematics, spelling learning and physics. We furthermore provide comparisons to previous student modeling approaches and analyze the influence of the different student modeling techniques on instructional policies. We demonstrate that our approach outperforms previous techniques in prediction accuracy on unseen data across all learning domains and yields meaningful instructional policies.

Index Terms—Bayesian networks, parameter learning, constrained optimization, error measures, instructional policies.

1 INTRODUCTION

Modeling and predicting student knowledge is a fundamental part of an intelligent tutoring system (ITS). In these systems, the selection of challenges to work on next is based on the estimation and prediction of the student’s current knowledge by the student model. The prediction accuracy and behavior of the student model directly influence the instructional policies [1] of the system and hence also the quality of teaching decisions. Therefore, an accurate student model is essential for individualization, i.e., the adaptation of the learning content and the difficulty level to the individual student. A large body of research was conducted in terms of modeling, i.e., the construction of accurate student models, and in terms of assessment, i.e., the definition of appropriate error measures for assessing prediction accuracy of student models.

Research in modeling covers a wide range of approaches for student modeling. Common techniques include logistic regression models and probabilistic models. Two of the most popular algorithms for estimating student knowledge are performance factors analysis (PFA) [2] and Bayesian Knowledge Tracing (BKT) as presented by Corbett and Anderson [3].

Traditional BKT and improved versions have been successfully used in different tutoring systems [4], [5]. An important task when using BKT is parameter learning. BKT models have been fit using brute-force grid search [6], gradient descent [7] or expectation maximization [8]. A large body of research was conducted to construct BKT models with higher prediction accuracy. Previous work includes clustering approaches [9] or individualization techniques, such as learning student- and skill-specific parameters [7], [10], [11] or modeling the parameters per school class [12].

Exhibiting a tree structure (directed acyclic variable dependencies with a single root variable), BKT allows for efficient parameter learning and accurate inference. However, tree-like models lack the ability to represent the hierarchy and relationships between the different skills of a learning domain. Employing dynamic Bayesian network models (DBN) has the potential to increase the representational power of the student model by explicitly modeling prerequisite skill hierarchies and hence further improve prediction accuracy. Furthermore, hierarchical models allow to incorporate expert knowledge and implicitly define the curriculum of the system. Indeed, hierarchical models have been successfully applied in ITS. Dynamic learning maps structure the sequence of topics hierarchically to reach certain thresholds in post-tests [13]. In cognitive item response theory, the attribute hierarchy method models hierarchically ordered attributes (competencies) required to correctly solve test problems [14]. Bayesian networks [15], [16], [17] have been applied to model skill dependencies at a single point in time.

DBNs, on the other hand, have been used to model and predict students’ performance [18], [19], engagement states [20], [21], and goals [18]. DBNs are also employed in user modeling [22]. In cognitive sciences, DBNs are applied to model human learning [23] and understanding [24]. Despite their beneficial properties to represent knowledge, DBNs have received less attention in modeling skill hierarchies as the joint representation of multiple skills over time in one network imposes challenges for learning and inference.
Lately, research in assessment of student models has gained increasing interest. Traditionally, the prediction accuracy of student models has been assessed using performance measures such as the root mean squared error (RMSE) [7], [12], [25], [26], the area under the ROC curve (AUC) [6], [27], [28], [29], [30], the mean absolute error (MAE) [31], [32] and the log-likelihood (LL) [2], [33]. Recently, the use of a single number as a performance measure has been questioned [34] and the different measures have been analyzed and discussed [34], [35]. Furthermore, new measures [34] and frameworks [36], [37] capturing model behavior in more detail have been introduced. Recent work has also analyzed the impact of the prediction accuracy on instructional policies [1].

In this paper, we contribute to both the research area of modeling and the research area of assessment: We present an efficient method for modeling prerequisite skill hierarchies and extensively evaluate our approach based on the latest research conducted in assessment of student models. We demonstrate how the topologies and relationships between different skills of a learning domain can be modeled using DBNs and show the advantages of using pre-requisite hierarchies for student modeling. Recently, [38] showed that a constrained latent structured prediction approach to parameter learning for DBNs yields accurate and interpretable models. Furthermore, DBNs modeling prerequisite skill hierarchies outperform traditional BKT regarding prediction accuracy [39]. We extend [39] by extensively evaluating the behavior of our hierarchical skill models regarding prediction accuracy and instructional policies and providing comparisons to the popular traditional approaches to student modeling. We define domain-specific DBNs modeling skill hierarchies for five large-scale data sets from different learning domains, containing up to 7000 students. To analyze the generalization capabilities of our approach, we have selected data sets generated from different tutoring systems addressing different fields of education such as mathematics, spelling learning and physics. Furthermore, students’ age ranges from elementary school to university level. Our results show that even simple skill hierarchies lead to significant improvements in prediction accuracy over BKT and PFA across all learning domains. Furthermore, the hierarchical skill models lead to meaningful instructional policies. By using the same constraints and parametrization for all our experiments, we also demonstrate that basic assumptions about learning hold across different learning domains.

## 2 Background: Student Models

In this section, we provide an overview of popular approaches for student modeling, which we later use for comparison in the experimental evaluation of our method.

### 2.1 Latent Factors Models

Student models applying logistic regression are based on the idea that the probability of a correct response to a task can be represented by a mathematical function of student and skill parameters. In such models, the binary task outcomes (correct/wrong) of the students follow a Bernoulli distribution, i.e., a binomial distribution with \( n = 1 \). Letting \( y_{st} \in \{0,1\} \) denote the response of student \( s \) to task \( t \), we obtain \( y_{st} \sim \text{binomial}(1, p_{st}) \).

**Additive Factors Model (AFM).** The AFM [40], [41] is a logistic regression model, which models the probability \( p_{st} \) of solving a task \( t \) correctly as a function of three parameters, resulting in

\[
p_{st} = \left(1 + \exp(-\theta_s + \sum q_{kt}(\beta_k + \gamma_k \cdot T_{sk}))\right)^{-1}.
\]

Hereby the random effect \( \theta_s \sim \mathcal{N}(0, \sigma^2) \) denotes the student proficiency and the fixed effects \( \beta_k \) and \( \gamma_k \) denote the difficulty and the learning rate for skill \( k \), respectively. \( T_{sk} \) is the number of practice opportunities student \( s \) has seen for skill \( k \) and \( q_{kt} \) is 1, if task \( t \) uses skill \( k \), and 0 otherwise.

**Performance Factors Analysis (PFA).** The PFA [2] is an extension of the AFM, with the following definition of the linear component:

\[
p_{st} = \left(1 + \exp(-\theta_s + \sum q_{kt}(\beta_k + \gamma_k \cdot S_{sk} + \rho_k \cdot F_{sk}))\right)^{-1}.
\]

In contrast to AFM, PFA differentiates between correct and incorrect practice opportunities: \( S_{sk} \) denotes the number of correctly solved tasks for student \( s \) at skill \( k \), while \( F_{sk} \) denotes the number of failures for student \( s \) at skill \( k \). The fixed effects \( \gamma_k \) and \( \rho_k \) therefore denote the learning rates for successes and failures, respectively.

### 2.2 Bayesian Knowledge Tracing

BKT models are a special case of DBNs [42] or more specifically of Hidden Markov Models (HMM), consisting of observed and latent variables. Latent variables represent student knowledge about one specific skill and are assumed to be binary, i.e., a skill can either be mastered by the student or not. The observed variables are also binary: they represent student answers (correct/incorrect) to questions associated with a specific skill and depend directly on the latent variables. The state of the latent variables is inferred based on these observations.

There are two types of parameters in an HMM: transition probabilities and emission probabilities. In BKT, the emission probabilities are defined by the slip probability \( p_s \) of making a mistake when applying a known skill and the guess probability \( p_G \) of correctly applying an unknown skill. The transition probabilities are described by the probability \( p_L \) of a skill transitioning from the unknown to the known state, while \( p_F \) is the probability of forgetting a previously known skill. In BKT, \( p_F \) is assumed to equal zero. The last parameter required to describe the BKT model is the initial probability \( p_0 \) of knowing a skill a-priori.

Employing one BKT model per skill, the learning task amounts to estimating the parameters given some observations. More specifically, given a sequence of observations \( y_m = (y_{m,1}, \ldots, y_{m,T}) \) with \( y_{m,t} \in \{0,1\} \) and time \( t \in \{1, \ldots, T\} \) for the \( m \)-th student with \( m \in \{1, \ldots, M\} \), what are the parameters \( \theta = \{p_0, p_L, p_F, p_S, p_G\} \) that maximize the likelihood \( \prod_m p(y_m | \theta) \) of the available data.

In BKT, this task is commonly solved using brute-force grid search [6], gradient descent [7] or expectation maximization [8].
3 METHODS: DYNAMIC BAYESIAN NETWORKS

DBNs have the potential to increase the representational power of the student model and therefore further increase prediction accuracy. Subsequently, we first give an introduction to DBNs using the traditional probabilistic notation, before introducing our log-linear formulation and the corresponding optimization. Finally, we also present a possible specification of a simple DBN.

When employing DBNs, we consider the different skills of a learning domain jointly within a single model. Student knowledge is again represented using binary latent variables (one per skill), which are updated based on observations associated with the skill under investigation. However, we now also model the dependencies between the different skills, e.g., two skills $S_A$ and $S_B$ are conditionally dependent if $S_A$ is a prerequisite for mastering $S_B$.

To illustrate this concept, consider Fig. 1 which shows an example DBN unrolled over $T$ time steps. The circular nodes represent binary skill variables: similar to BKT, a skill can be either mastered or not. The example network consists of three skills: $S_A$, $S_G$, and $S_N$. The rectangular nodes represent observable variables, i.e. binary student answers (correct/incorrect) to tasks associated with the respective skill. The index numbers indicate the time step.

3.1 Probabilistic Notation

The learning task of a DBN model is described as follows: let the set of $N$ variables of the model be denoted by $X = \{X_i \mid i \in \{1, \ldots, N\}\}$. The set of variables $X$ contains all skill nodes $S$ as well as all observation nodes $O$ of the model. In addition, let $H$ denote the domain of the unobserved variables, i.e., missing student answers and the binary skill variables, while $Y$ refers to the observed space, disjoint from the latent space $H$. Let us consider an example student $m$ for the model in Fig. 1 solving a task associated with skill $S_{G,1}$ in the first time step correctly, i.e., $O_{G,1} = 1$. The set of observed variables $y_m$ for the first time step consists of the student’s answer, i.e., $y_m = \{O_{G,1}\}$. The set of hidden variables $h_m$ for the first time step contains all other variables, i.e., $h_m = \{S_{A,1}, S_{N,1}, S_{G,1}, O_{N,1}\}$. During learning, we are interested in finding the parameters $\theta$ that maximize the likelihood of the observed data $y_m$ with $y_m = (y_{m,1}, \ldots, y_{m,T})$ representing a sequence of $T$ binary answers from the $m$-th student, while $h_m$ denotes the hidden variables for student $m$. The log-likelihood of a DBN [43] is then given by

$$L(\theta) = \sum_m \ln \left( \sum_{h_m} p(y_m, h_m \mid \theta) \right), \quad (3)$$

where we marginalize over the states of the hidden variables $h_m$ for student $m$. The joint probability $p(y_m, h_m \mid \theta)$ of the model for student $m$ is defined via

$$p(y_m, h_m \mid \theta) = \prod_i p(X_{m,i} = x_{m,i} \mid pa(X_{m,i}) = x_{m,pa(X_{m,i})}) = \prod_i p_{ijm, km,i}, \quad (4)$$

where $x_{m,i}$ is the $i$-th variable in the model for student $m$ and $pa(X_{m,i})$ are the parents of $X_{m,i}$ while $x_{m,i}$ and $x_{m,pa(X_{m,i})}$ are the realizations of the random variables $X_{m,i}$ and $pa(X_{m,i})$, i.e., the states assigned to $X_{m,i}$ and $pa(X_{m,i})$ given by $(y_m, h_m)$. Furthermore, we let $j_{m,i} := x_{m,i}$ and $k_{m,i} := x_{m,pa(X_{m,i})}$ to simplify the notation. Therefore, $p_{ijm, km,i}$ denotes exactly one entry in the conditional probability table (CPT) of $X_{m,i}$.

3.2 Log-linear models

The log-likelihood of a DBN can alternatively be formulated using a log-linear model. This formulation is flexible and predominantly used in literature [44], [45]. Therefore, we reformulate the learning task in the following. Let $\phi : \mathcal{Y} \times \mathcal{H} \rightarrow \mathbb{R}^F$ denote a mapping from the latent space $\mathcal{H}$ and the observed space $\mathcal{Y}$ to an $F$-dimensional feature vector. The log-likelihood from Eq. (3) can then be reformulated to

$$L(w) = \sum_m \ln \left( \sum_{h_m} \exp \left( w^T \phi(y_m, h_m) - \ln(Z) \right) \right), \quad (5)$$

where $Z$ is a normalizing constant and $w$ denotes the weights of the model. Next, we show that this log-linear formulation of the log-likelihood is equivalent to the traditional notation. Comparing Eq. (5) to Eq. (3), it follows that

$$\prod_i p_{ijm, km,i} = \frac{1}{Z} \exp w^T \phi(y_m, h_m) = \frac{1}{Z} \exp \sum_i w_i^T \phi_i(y_m, h_m), \quad (6)$$

and therefore

$$\forall i, j, k : p_{ijk} = \frac{1}{Z} \exp w_i^T \phi_i(x), \quad (7)$$

where $x$ are the realizations of all random variables in $X$ with $j \in x$ and $k \subset x$. A feature vector $\phi$ and weights $w$ that fulfill Eq. (7) can be specified as follows: consider the CPT describing the relationship between a node $X_A$ and its $n - 1$ parent nodes $pa(X_A)$. The CPT for these $n$ nodes contains $2^n$ entries. Let $k \in \{0, 1\}^{n-1}$ denote one possible assignment of states to the parent nodes $pa(X_A)$.
We can therefore define \( p(X_A = 1 \mid p_a(X_A) = k) = 1 - p(X_A = 0 \mid p_a(X_A) = k) = 1 - p_{A|D,k} \). To continue, let \( p_{A|x,k} = \frac{1}{2} \exp w_{A,k}(1 - 2x_A) = \exp w_{A,k}(1 - 2x_A) / (\exp w_{A,k} + \exp(-w_{A,k})) \), which leads to the feature function \( \phi_A(x) = 1 - 2x_A \). We therefore obtain the joint distribution as a product of the exponential terms which translates to a weighted linear combination of feature vector entries in the exponent and thus fulfills Eq. (7). From this formulation also follows that we need \( 2^n - 1 \) parameters to specify a CPT including \( n \) skills.

### 3.3 Optimization

Due to the loopy structure, the learning task for DBNs is computationally more expensive than the one for HMMs. However, [45] showed that some approximations admit efficient parameter learning. Note that interpretability of the parameters is not ensured, since guarantees exist only for converging to a local optimum, i.e., the approximation might result in degenerate models with for example a probability of guessing \( p_G > 0.5 \). Therefore, [38] extended the approach presented by [45] to include constraints on parameters and demonstrated that the constrained optimization increases prediction accuracy on unseen data while guaranteeing interpretable models. Using the log-linear formulation, the algorithm presented in [38] can be directly applied to learn the parameters of a DBN model.

### 3.4 Specification

Next, we explain the specification of a simple DBN. The example DBN illustrated in Fig. 1 consists of three skills: \( S_A, S_G, \) and \( S_N \). Two of the skills (\( S_G \) and \( S_N \)) have tasks associated with them, while the third skill \( S_A \) cannot be observed. Similarly to BKT, we can interpret the parameters of a DBN in terms of a learning context. To specify the CPTs of the example DBN, we employ \( F = 22 \) weights that can be associated with a parameter set \( \theta \). We subsequently use \( \sim \) to denote proportionality in the log domain; i.e., \( w \sim p \) is equivalent to \( w \propto \exp p \). Let \( O_N \) denote the task associated with skill \( S_N \). Then the parameters \( w_{00} \approx p(O_N = 0 \mid S_N = 0) = 1 - p_G \) and \( w_{25} \approx p(O_N = 0 \mid S_N = 1) = p_S \) represent the guess and slip probabilities. Similarly, \( w_{28} \) and \( w_{19} \) are associated with \( p_G \) and \( p_S \) as evident from Fig. 1. Furthermore, parameters \( w_{6} \approx p(S_{A,t} = 0 \mid S_{A,t-1} = 0) = 1 - p_L \) and \( w_{7} \approx p(S_{A,t} = 0 \mid S_{A,t-1} = 1) = p_F \) are associated with learning and forgetting; the same holds true for \( w_8 \) and \( w_9 \).

Skills \( S_A \) and \( S_G \) are prerequisites for knowing skill \( S_N \), i.e., the probability that skill \( S_N \) is mastered in time step \( t \) depends not only on the state of skill \( S_N \) in the previous time step, but also on the states of \( S_A \) and \( S_G \) in the current time step. Therefore \( w_{10} \approx p(S_{N,t} = 0 \mid S_{N,t-1} = 0, S_{A,t} = 0, S_{G,t} = 0) = 1 - p_{L0} \), where \( p_{L0} \) denotes the probability that the student learns \( S_N \) despite not knowing \( S_A \) and \( S_G \). Also, \( w_{17} \approx p(S_{N,t} = 0 \mid S_{N,t-1} = 1, S_{A,t} = 1, S_{G,t} = 1) = p_{F1} \), the probability of forgetting a previously learnt skill. Furthermore, we set \( w_{1} \approx 1 - p_{LM} \) if \( l \in \{11, 12, 13\} \) and \( w_{7} \approx 1 - p_{FM} \) if \( l \in \{14, 15, 16\} \), where \( p_{LM} \) denotes the probability that the student learns \( S_N \) given that he knows at least one of the precursor skills of \( S_N \). Moreover, \( p_{FM} \) is the probability that the student forgets the previously known skill \( S_N \), when either \( S_A \) or \( S_G \) or none of them are known.

Finally, the parameters \( w_l \) with \( l \in \{2, 3, 4, 5\} \) describe the dependencies between the different skills. We let \( w_{l} \approx 1 - p_{p_0} \), if \( l \in \{2, 3, 4\} \) and \( w_{5} \approx p_{p_1} \), where \( p_{p_0} \) is the probability of knowing a skill despite having mastered only part of the prerequisite skills and \( p_{p_1} \) denotes the probability of failing a skill given that all precursor skills have been mastered already. Moreover, we refer to the probability of knowing a skill a-priori via \( p_0 \). Note that \( w_0 \) and \( w_1 \) are associated with \( p_0 \). The example DBN can therefore be described by the parameter set \( \theta = \{p_0, p_G, p_F, p_L, p_F_1, p_M, p_F_3, p_F_6, p_F_9, p_F_11, p_F_13, p_F_15, p_F_17, p_F_19\} \). Importantly, the method proposed in this work is independent of the exact parameterization used. The parameterization introduced here serves as an example only and could be easily extended.

### 4 Experimental Evaluation

We show the benefits of DBN models with higher representational power on five data sets from various learning domains. The data sets were collected with different tutoring systems and contain data from elementary school students up to university students. With the extensive experimental evaluation, we aim at answering the following three research questions. 1) What are the benefits in terms of prediction accuracy of DBNs representing skill hierarchies compared to traditional student modeling approaches assuming independence between the different skills? 2) Where do the differences in prediction accuracy between the different student models come from? 3) How do the different student modeling techniques influence instructional policies?

#### 4.1 Experimental Setup

Evaluation of the models was performed using a training-test setting: For all the different models, we learned the parameters on a training set and evaluated their performance on a test set. All evaluation measures were computed using student-stratified 10-fold cross validation.

Fitting the BKT models was done using [7], applying skill-specific parameters and using gradient descent for optimization. As described in [7], we set the forget probability \( p_F \) to 0, while \( p_S \) and \( p_G \) are bounded by 0.3. In the following, we will denote this constrained BKT version as \( BKT_C \). Since parameter fitting for BKT models converges to different values depending on the initialization of the algorithm [6], [28], [46], we generated ten different \( BKT_C \) models using different start values. We included the default values provided by [7] as well as nine other randomly generated sets of start values fulfilling the constraints. Subsequently, we selected the model with the best prediction accuracy in terms of root mean squared error on the test set for comparison. This selection process overestimates the performance of BKT and therefore provides an upper bound on the prediction accuracy of BKT.

The parameters of the latent factors models were trained using the lme4 package of R. AFM and PFA require a student parameter (the student proficiency \( \theta_s \)). For the unseen students in the test sets, we set \( \theta_s \) to the mean of the trained student parameters.
We used constrained latent structured prediction [38] to learn the parameters of the DBNs. All models were parametrized according to Sec. 3.4 and we imposed the constraints described in the following on the parameter set \( \theta \) of the different models to ensure interpretable parameters. For our first constraint set \( C_1 \), we let \( p_D \leq 0.3 \) for \( D \in \{ G, S, L, F, L0, F1 \} \) to ensure that parameters associated with guessing, slipping, learning and forgetting remained plausible. The constraints on \( \theta \) can be directly turned into constraints on \( w \). For the example DBN (Fig. 1), the constraints translate into the following linear constraints on the weights for \( C_1 \): \( w_i \geq 0.4236 \), if \( i \in \{ 6, 8, 10, 18, 20 \} \) and \( w_i \leq -0.4236 \), if \( i \in \{ 7, 9, 17, 19, 21 \} \). For the second constraint set \( C_2 \), we augmented \( C_1 \) by limiting \( p_D \leq 0.3 \) if \( D \in \{ LM, FM, P0, P1 \} \), yielding \( w_i \geq 0.4236 \), if \( i \in \{ 2, 3, 4, 11, 12, 13 \} \) and \( w_i \leq -0.4236 \), if \( i \in \{ 5, 14, 15, 16 \} \) for the example DBN (Fig. 1). The additional constraints ensure that parameters are consistent with the hierarchy assumptions of the model. The constraint sets \( C_3 \) and \( C_4 \) bound the same parameters as \( C_1 \) and \( C_2 \), but are more restrictive by replacing 0.3 by 0.2. Note that for these examples constraints were selected according to previous work [3]. Similarly to BKT, we ran the optimization ten times for each constraint set using different starting values. Subsequently, we selected the model with the highest likelihood on the training data set for comparison.

Prediction for the probabilistic models was performed as follows: we assumed the observation at time \( t = 1 \) to be given and predicted the outcome at time \( t \) with \( t \in \{ 2, \ldots, T \} \) based on the previous \( t - 1 \) observations. To predict the outcome of the latent factors models, we evaluated the trained regression model at time \( t \). The number of observations \( T \) for the different experiments is the minimal number of observations, which includes observations of all skills of the according model.

### 4.2 Data Sets and Models

We evaluated our DBN models on five large-scale data sets. To analyze the generalization capabilities of our approach, we selected data sets according to the following criteria: the data logs come from different ITS (Calcularis, Andes2, Cognitive Tutor, and Dybuster) and cover a wide range of age classes (from elementary school children over high school children up to university students) as well as different learning domains (mathematics, physics, spelling). In the following, we describe the five data sets as well as their respective DBNs. Note that we use the same skill sets for all our models, i.e., the skill sets described for the DBNs are also used for BKT, AFM, and PFM. We further assume that each observation is associated with exactly one skill.

**Number Representation.** To build the first model, we used data collected from Calcularis, an ITS for elementary school children with math learning difficulties [47]. Calcularis turns current neuro-cognitive theory into the design of different instructional games training number representations and number understanding as well as arithmetic operations. Student knowledge is represented as a DBN consisting of different mathematical skills [48], [49].

The data set contains log files of 1581 children with at least five sessions of 20 minutes per user.

The graphical model used in this experiment (see Fig. 1) is an excerpt of the skill model of Calcularis described in [48]. Skill \( S_A \) represents knowledge about the Arabic notation system. Calcularis does not contain any tasks associated with this skill. The ability to assign a number to an interval is denoted by \( S_G \). The task associated with this skill is to guess a number in as few steps as possible. Finally, \( S_N \) denotes the ability to indicate the position of a number on a number line. We used a maximum of \( T = 100 \) observations per child for learning and prediction and specified the CPTs of the graphical model with \( F = 22 \) weights as illustrated in Fig. 1.

**Subtraction.** The second model is based on the same data set as the first model. This time, however, we investigated subtraction and number understanding skills. The graphical model (see Fig. 2) is again an excerpt of the skill model [48] of Calcularis. Subtraction skills are ordered according to their difficulty, which is determined by the magnitude of involved numbers, task complexity and the means allowed to solve a task. Skills \( S_1 \) (e.g., 48-6=?), \( S_2 \) (e.g., 48-9=?), \( S_3 \) (e.g., 48-26=?), \( S_4 \) (e.g., 48-29=?), and \( S_5 \) denote subtraction tasks in the number range 0–100. We emphasize that there are no observation nodes associated with \( S_1 \) and \( S_5 \). The number understanding skill \( S_6 \) represents knowledge about the relational aspect of number (i.e., a number can be seen as a difference between two other numbers) in the number range 0–1000. Finally, skills \( S_7 \) (e.g., 158-3=?), \( S_8 \) (e.g., 158-3=?), \( S_9 \) (e.g., 158-9=?), and \( S_3 \) (e.g., 158-9=?), \( S_5 \) and \( S_6 \) denote subtraction tasks in the number range 0–1000. The difference between \( S_7 \) and \( S_9 \) lies in the means allowed to solve the task. A maximum of \( T = 100 \) observations per child was used for learning and prediction. Specification of the CPTs for the model requires \( F = 86 \) weights as illustrated in Fig. 2.

**Physics.** To build the third model, we used the ‘USNA Physics Fall 2005’ data set accessed via DataShop [50]. Data originate from 77 students of the United States Naval Academy and were collected from Andes2, an ITS for physics [18]. The tutor uses rule-based algorithms to build solution graphs that identify all possible solutions to the different tasks. Based on these graphs, a Bayesian network is constructed to assess the general physics knowledge of the student as well as the progress for the problem at hand. We used the different modules of the data set as skills for our experiment. The graphical model is depicted in Fig. 3. Note that, since we are neither experts regarding physics education nor concerning Andes2, we intentionally used a simplified model representing a subset of the modules as skills to avoid introducing incorrect assumptions. The model consists of the following modules: “Vectors” (\( S_V \)), “Translational Kinematics” (\( S_K \)), “Statistics” (\( S_S \)) and “Dynamics” (\( S_D \)). These modules consist of complex tasks for the given topic, i.e., calculating total forces in a system (see example in [18]). A maximum of \( T = 500 \) observations per child were considered for learning and prediction and the model can be described by \( F = 33 \) weights.
Algebra. This model is based on data from the KDD Cup 2010 Educational Data Mining Challenge (http://pslcdatashop.web.cmu.edu/KDDCup). The data set contains log files of 6043 students that were collected by the Cognitive Tutor [4], an ITS for mathematics learning. The student model applied in this system is based on BKT, since the cognitive theory behind the cognitive tutor assumes skills to be independent.

We used the units of the ‘Bridge to Algebra’ course as skills for our experiment and selected four units of increasing difficulty, where students have to solve word problems involving calculations with whole numbers. The graphical model for this experiment is illustrated in Fig. 4. Skill $S_A$ (e.g., 728624 – 701312) denotes written addition and subtraction tasks without carrying/borrowing, while $S_S$ involves carrying/borrowing (e.g., 728624 – 703303). $S_M$ (e.g., 33564 × 18) and $S_D$ (e.g., 10810 ÷ 46) represent long multiplications and divisions. The original model for the data set used contains 94 skills that were carefully designed by domain experts. Building a fine-grained skill hierarchy would also require domain expert advice. Therefore, we again employed a simplified skill model representing a subset of the modules as skills. We used a maximum of $T = 500$ observations per student (consisting only of ‘correct on first attempt responses’) for learning and prediction and specified the CPTs of the model employing $F = 29$ weights.

Spelling. The last model uses data collected from Dybuster, an ITS for elementary school children with dyslexia [51]. The data set at hand contains data of 7265 German-speaking children. Dybuster groups the words of a language into hierarchically ordered modules with respect to their frequency of occurrence in the language corpus as well as a word difficulty measure. The latter is computed based on the word length, the number of dyslexic pitfalls and the number of silent letters contained in the word.

We used these modules as skills to build our graphical model (see Fig. 5). Skills $S_E$, $S_M$ and $S_D$ denote the modules 3, 4 and 5 within Dybuster. Word examples are “warum” (“why”, $S_E$), “Donnerstag” (“Thursday”, $S_M$) and “Klapperschlange” (“rattlesnake”, $S_D$). We used a maximum of $T = 200$ observations per child for the learning and prediction tasks and parametrized the model using $F = 21$ weights.
4.3 Prediction Accuracy

To answer our first research question What are the benefits of prediction accuracy of DBNs representing skill hierarchies compared to traditional student modeling approaches assuming independence between the different skills?, we assessed the predictive performance of the different models using the root mean squared error (RMSE) and the area under the ROC curve (AUC). The RMSE is widely used in educational data mining to measure the performance of skill models and has demonstrated a high correlation to the ‘moment of learning’ (step in which a student learnt a skill) for BKT models [35], [52]. The AUC is also often applied for the evaluation of educational models [6], [27], [28], [29], [30], but its use as the only performance metric has been questioned lately [34]. However, in our case the AUC is used as an additional measure to the RMSE to assess the models’ abilities to discriminate failures from successes. All error measures were calculated using 10-fold student-stratified cross validation. Statistical significance was computed using a two-sided t-test, correcting for multiple comparisons (Bonferroni-Holm). Tab. 1 details the RMSE and AUC for all models and data sets. Bold numbers denote a significant improvement of the constrained DBN model (for the different constraints $C_1 - C_4$ on the probabilities described in Sec. 4.1) over all alternative models (BKT$_C$, AFM and PFA). The best performing model for each measure is marked (*).

The constrained DBN approach yields significant and large improvements in prediction accuracy compared to BKT$_C$ for the Number Representation data set. We highlight the reduction of the RMSE by 3.8% and the large improvement achieved in AUC (AUC$_{BKT_C} = 0.5995$, AUC$_{C_4} = 0.7026$). The constrained DBN models also outperform the logistic regression models in terms of RMSE and AUC with a reduction of the RMSE of 5.9% for AFM and 4.5% for PFA. AFM performance worst for both measures on this data set. While PFA exhibits an RMSE similar to BKT$_C$, it shows a far better AUC than BKT$_C$ (AUC$_{BKT_C} = 0.5995$, AUC$_{PFA} = 0.6717$). Note that all constrained DBN models significantly outperform all alternative models on both measures.

The resulting prediction accuracy for the Subtraction data set again demonstrates that the hierarchical DBN model outperforms the BKT$_C$ model as well as the logistic regression models. With a reduction of the RMSE by 3.5% compared to BKT$_C$ and a reduction in RMSE of 3.0% over the PFA model, the benefits of the DBN model are again substantial. The DBN model exhibits a large improvement in AUC over BKT$_C$ (AUC$_{BKT_C} = 0.5990$, AUC$_{C_4} = 0.6928$). Note that the PFA model again significantly outperforms BKT$_C$ regarding the AUC. The DBN model shows a large growth in AUC (compared to PFA) for the more restrictive constraint sets (AUC$_{C_4} = 0.6882$, AUC$_{C_3} = 0.6928$, AUC$_{PF_A} = 0.6532$).
For the Physics data set, we observe significant and substantial improvements of the AUC over all alternative models (AUC\(_{C_3} = 0.7021\), AUC\(_{BKT_C} = 0.4991\), AUC\(_{PFA} = 0.5807\), AUC\(_{AFM} = 0.5425\)). Furthermore, the RMSE is reduced by 6.3% compared to BKT\(_C\) and by 6.6% compared to PFA. For this data set, the logistic regression models exhibit an RMSE similar to BKT\(_C\), however, they show a substantially larger AUC than BKT\(_C\).

On the Algebra data set, all DBN model configurations significantly outperform all other models regarding the RMSE and the AUC. We observe an improvement in RMSE of 3.7% compared to BKT\(_C\) and of 3.8% and 5.1% compared to PFA and AFM, respectively. The lower RMSE of all the models compared to the other data sets probably stems from the bias towards correct observations, i.e., this data set contains 85% of correct observations. Improvements in AUC are also substantial with all models exhibiting much lower values than DBN (AUC\(_{C_2} = 0.7042\), AUC\(_{BKT_C} = 0.6012\), AUC\(_{AFM} = 0.6034\), AUC\(_{PFA} = 0.6407\)).

Performance differences between the four models tend to be low for the Spelling data set. While still being significant, the absolute improvement in RMSE of the DBN models compared to BKT\(_C\) drops to 0.1%. The logistic regression models show a similar performance (\(\text{RMSE}_{PFA} = 0.4481\), \(\text{RMSE}_{AFM} = 0.4498\)). The DBN models again outperform AFM and BKT\(_C\) regarding the AUC (AUC\(_{C_2} = 0.5808\), AUC\(_{BKT_C} = 0.5029\), AUC\(_{PFA} = 0.5495\)) and perform in range with PFA (AUC\(_{PFA} = 0.5790\)). However, AUC values for all the models tend to be lower than the values achieved on the other data sets.

4.4 Model Behavior

Our second research question Where do the differences in prediction accuracy between the different student models come from? aims at understanding model behavior in more detail and investigating the observed differences in prediction accuracy (see Sec. 4.3) between the different approaches. To answer this question, we computed the Brier score of the different models [34]. The Brier score is equivalent to the mean squared error (MSE) and can be decomposed into three components: BS = REL - RES + UNC. The uncertainty (UNC) quantifies the inherent uncertainty of the events (correct and incorrect outcomes) in the data set and is independent of the selected model. The reliability term (REL) measures the difference between predicted and observed probabilities. A reliability of zero (REL = 0) indicates a perfectly reliable prediction. The resolution term (RES) measures how much predictions differ from the base rate (proportion of correct outcomes in the data set) and therefore gives and indication of the range of predicted probabilities, i.e., a high resolution is desirable. An ideal model would therefore minimize the REL term, while maximizing the RES term. We computed the Brier score and its components using 10-fold student-stratified cross validation. Statistical significance was calculated using a two-sided t-test, correcting for multiple comparisons (Bonferroni-Holm). The reliability and resolution for the different data sets are displayed in Fig. 6. For the DBNs, we included only the model with the best RMSE for each data set (see Tab. 1).

For the Number Representation data set, BKT\(_C\) exhibits a significantly better reliability than the DBN model with the best RMSE (REL\(_{BKT_C} = 0.0732\), REL\(_{C_2} = 0.0859\)). The reliabilities of the DBN model and the PFA model are in range (REL\(_{PFA} = 0.0929\), while AFM shows a much worse reliability (REL\(_{AFM} = 0.1178\)), but has the best resolution (RES\(_{AFM} = 0.0919\)). Resolutions of the DBN and PFA models are again similar (RES\(_{PFA} = 0.0737\), RES\(_{C_2} = 0.0685\)). For this data set, BKT\(_C\) achieves a significantly lower resolution than DBN (RES\(_{BKT_C} = 0.0332\)), which explains the performance differences in RMSE.

We observe a different picture for the Subtraction data set. While the DBN and PFA models show again similar reliabilities (REL\(_{C_2} = 0.0522\), REL\(_{PFA} = 0.0557\)), BKT\(_C\) exhibits a worse reliability (REL\(_{BKT_C} = 0.0671\)); however, the differences in reliability between BKT\(_C\) and DBN are not statistically significant. The DBN model again demonstrates a significantly higher resolution than BKT\(_C\) (RES\(_{C_2} = 0.0406\), RES\(_{BKT_C} = 0.0241\)). Also the results for the AFM model are confirmed, which shows the worst reliability (REL\(_{AFM} = 0.0741\)), but a high resolution (RES\(_{AFM} = 0.0445\)).

On the Physics data set, the BKT\(_C\), PFA, and DBN models shows similar reliabilities (REL\(_{C_2} = 0.0127\), REL\(_{PFA} = 0.0141\), REL\(_{BKT_C} = 0.0174\)), the AFM again demonstrates a significantly worse reliability than the DBN model (REL\(_{AFM} = 0.0237\)). For this data set, the DBN model exhibits a significantly higher resolution than all the other models (REL\(_{C_2} = 0.0262\), RES\(_{BKT_C} = 0.0016\), RES\(_{PFA} = 0.0096\), RES\(_{AFM} = 0.0138\)).

The DBN model achieves a better reliability than all the other models on the Algebra data set, with all differences being statistically significant (REL\(_{C_2} = 0.0170\), REL\(_{PFA} = 0.0246\), REL\(_{BKT_C} = 0.0216\), REL\(_{AFM} = 0.0275\)). The same fact holds true for the resolution, where the DBN model again significantly outperforms all the other models (RES\(_{C_2} = 0.0160\), RES\(_{BKT_C} = 0.0009\), RES\(_{PFA} = 0.0127\), RES\(_{AFM} = 0.0152\)). This result explains the lower RMSE achieved for all tested DBN models (see Tab. 1).

All models achieve similar reliability on the Spelling data set (REL\(_{C_2} = 0.0284\), REL\(_{PFA} = 0.0262\), REL\(_{BKT_C} = 0.0288\), REL\(_{AFM} = 0.0294\)), with all the small differences between DBN and the other models being statistically significant. There is no significant difference between the resolution of the DBN and the PFA models (RES\(_{C_2} = 0.0160\), RES\(_{BKT_C} = 0.0009\), RES\(_{PFA} = 0.0127\), RES\(_{AFM} = 0.0152\)). This result explains the lower RMSE achieved for all tested DBN models (see Tab. 1).

4.5 Instructional Policies

Lately, the influence of small differences in performance measures on the learning outcome has been questioned [25]. However, it has been shown that models with similar predictive performance can lead to significantly different instructional policies [1], [36], which leads to our third research question: How do the different student modeling techniques influence instructional policies? To evaluate the impact of the different models on the task selection, we calculated the effort \(E\) (number of practice opportunities needed to pass a skill) and score \(S\) (percentage of correct observations after having the skill passed) of the different models. We used the Teal framework [36] to define the mastery criterion:
TABLE 1

<table>
<thead>
<tr>
<th>Number Representation</th>
<th>RMSE</th>
<th>DBN_{C_1}</th>
<th>DBN_{C_2}</th>
<th>DBN_{C_3}</th>
<th>DBN_{C_4}</th>
<th>BKT_C</th>
<th>AFM</th>
<th>PFA</th>
</tr>
</thead>
<tbody>
<tr>
<td>RMSE</td>
<td>0.4469</td>
<td>0.4452</td>
<td>0.4416</td>
<td>0.4378*</td>
<td>0.4350</td>
<td>0.4657</td>
<td>0.4586</td>
<td></td>
</tr>
<tr>
<td>AUC</td>
<td>0.7008</td>
<td>0.7026*</td>
<td>0.7010</td>
<td>0.6971</td>
<td>0.5995</td>
<td>0.6262</td>
<td>0.6717</td>
<td></td>
</tr>
<tr>
<td>Subtraction</td>
<td>RMSE</td>
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<td>0.4215*</td>
<td>0.4309</td>
<td>0.4216*</td>
<td>0.4368</td>
<td>0.4457</td>
<td>0.4347</td>
</tr>
<tr>
<td></td>
<td>AUC</td>
<td>0.6166</td>
<td>0.6882</td>
<td>0.6344</td>
<td>0.6928*</td>
<td>0.5990</td>
<td>0.5763</td>
<td>0.6532</td>
</tr>
<tr>
<td>Physics</td>
<td>RMSE</td>
<td>0.4521</td>
<td>0.4272</td>
<td>0.4244*</td>
<td>0.4465</td>
<td>0.4530</td>
<td>0.4527</td>
<td>0.4497</td>
</tr>
<tr>
<td></td>
<td>AUC</td>
<td>0.6444</td>
<td>0.7009</td>
<td>0.7021*</td>
<td>0.6874</td>
<td>0.4991</td>
<td>0.5425</td>
<td>0.5807</td>
</tr>
<tr>
<td>Algebra</td>
<td>RMSE</td>
<td>0.3325</td>
<td>0.3246*</td>
<td>0.3311</td>
<td>0.3260</td>
<td>0.3369</td>
<td>0.3419</td>
<td>0.3374</td>
</tr>
<tr>
<td></td>
<td>AUC</td>
<td>0.6711</td>
<td>0.7042*</td>
<td>0.6731</td>
<td>0.7040</td>
<td>0.6012</td>
<td>0.6034</td>
<td>0.6407</td>
</tr>
<tr>
<td>Spelling</td>
<td>RMSE</td>
<td>0.4521</td>
<td>0.4495</td>
<td>0.4491</td>
<td>0.4472*</td>
<td>0.4503</td>
<td>0.4498</td>
<td>0.4481</td>
</tr>
<tr>
<td></td>
<td>AUC</td>
<td>0.5699</td>
<td>0.5775</td>
<td>0.5737</td>
<td>0.5808*</td>
<td>0.5029</td>
<td>0.5495</td>
<td>0.5790</td>
</tr>
</tbody>
</table>

Fig. 6. Reliability (REL) and Resolution (RES) over the different data sets for BKT_C, AFM, PFA and the DBN model with the lowest RMSE.

Instruction for a skill \( s \) is stopped, when the predicted probability \( \hat{p}_t \) of correctly answering the next task \( t \) associated with skill \( s \) is above a given threshold \( R \), i.e., \( \hat{p}_t > R \). Since the effort and score are dependent on the threshold \( R \) employed, we computed the measures for a range of thresholds \( R \in [0.6,...,0.8] \). We again calculated all measures using 10-fold student-stratified cross validation. We included only one logistic regression model (PFA) in this evaluation, since PFA consistently outperformed AFM on the performance measures over all data sets. Furthermore, we again only included the DBN model with the best RMSE (see Tab. 1). Fig. 7 shows the resulting effort (blue) and score (green) for DBN, BKT_C and PFA at different thresholds \( R \) over all data sets.

On the Number Representation data set, the three models show similar scores. The maximum score (\( S = 0.85 \)) for the DBN model is reached at \( R = 0.75 \) with an average effort \( E = 14.92 \). An effort of \( E = 14.92 \) can be considered as large, since it meets the wheel-spinning criterion (\( E > 10 \) [53]). However, the model still reaches a score \( S = 0.82 \) with an average effort below this limit (\( E = 9.4 \)). The BKT_C model achieves the lowest maximum score \( S = 0.79 \) of all the models, however, the average effort \( E = 3.60 \) is also low. Note that effort and score for the BKT_C model could be computed only for \( R \leq 0.69 \) since the model never predicted a probability \( \hat{p}_t > 0.69 \). The values for the scores show very little differences, this behavior can be explained by the low resolution of the BKT_C model (see Fig. 6, bottom) on this data set. The PFA model exhibits the maximum score (\( S = 0.83 \)) at \( R = 0.66 \) with the highest average effort of \( E = 17.90 \). The PFA model generally predicts flatter learning curves on this data set, resulting in the higher effort \( E \) needed for the given thresholds.

For the Subtraction data set, the results look similar. The BKT_C model exhibits a flat score curve (maximum score \( S = 0.68 \), but also the lowest average effort (\( E = 6.00 \). The DBN model again achieves a considerably higher score (\( S = 0.81 \)) than the BKT_C model with a reasonable effort \( E = 10.07 \). The PFA model predicts student performance much more conservatively than the other models, resulting in high efforts necessary to reach \( \hat{p}_t > R \). The PFA model achieves scores around \( S = 0.8 \) for an average effort \( E > 15 \).

On the Physics data set, effort (\( E = 3.00 \)) and score (\( S = 0.70 \)) take just one value for the BKT_C model. This is due to the small resolution (\( RES_{BKT_C} = 0.0016 \)) of the BKT_C model on this data set (illustrated in Fig. 6, bottom). For the PFA model, scores increase with a higher effort, but the model generally exhibits very high efforts. However, for an effort \( E = 3.00 \), the model achieves a score \( S = 0.71 \), which is in the range of the BKT_C model. Also the DBN model achieves a score \( S = 0.70 \) with an effort \( E = 3.21 \) on average. Similarly to the BKT_C model, scores do not increase with higher effort.

For the Algebra data set, all models achieve higher scores (\( S_{BKT_C} = 0.85, S_{PFA} = 0.85, S_{AFM} = 0.85 \)) than on the other data sets with similar efforts (\( E_{BKT_C} = 4.00, E_{PFA} = 4.00, E_{AFM} = 5.75 \)). This observation is probably due to the fact that the Algebra data set contains easier skills; this bias towards correct observations is also visible in the low RMSE of all models for this data set (see Tab. 1). For the PFA model, scores get as high as \( S = 0.90 \), but also the average effort grows to \( E = 349.27 \). This high number of observations is possible due to the simplified skill model: we used the units...
of the course as skills to build the model.

The effort and score for the different models on the Spelling data set confirm the previous results: We observe similar scores ($S_{\text{BKT}_C} = 0.71$, $S_{\text{PFA}} = 0.70$, $S_{\text{C}} = 0.71$) with reasonable efforts ($E_{\text{BKT}_C} = 3.00$, $E_{\text{PFA}} = 3.00$, $E_{\text{C}} = 6.00$) for the three models. The BKT$_C$ model again takes only one value for effort and score. For the DBN model, the scores do not increase with ascending effort. The scores for the PFA model slowly increase over time, however, they drop for very high efforts.

5 Discussion

The student model is an integral part of an ITS. One of the most popular approaches to student modeling is BKT. However, this technique cannot model the relationships between the different skills of a learning domain. More complex DBN models have the potential to increase the representational power by modeling skill hierarchies. In this work, we have suggested the use of DBNs representing skill hierarchies for student modeling and have extensively evaluated them on five large-scale data sets. Our experiments answer three important research questions.

The first research question analyzed the prediction accuracy of our approach: What are the benefits in terms of prediction accuracy of DBNs representing skill hierarchies compared to traditional student modeling approaches assuming independence between the different skills?. Our results demonstrate that DBN models outperform BKT in prediction accuracy. For hierarchical domains, the RMSE can be reduced by 3.5% (Subtraction data set) to 6.3% (Physics data set). As expected, adding skill topologies has a much smaller benefit for learning domains that are less hierarchical in nature (such as spelling learning). While differences in RMSE are still statistically significant, the magnitude of improvement is low. The results obtained on the Physics and Algebra data sets show that even simple hierarchical models improve prediction accuracy significantly. A domain expert employing a more detailed skill topology and more complex constraint sets could probably obtain an even higher accuracy on these data sets. Our comparison to logistic regression models also shows the superiority of the DBN models: The DBN models significantly outperform PFA on all data sets. PFA exhibits a similar or better RMSE than BKT$_C$ on the data sets. These results are not unexpected since PFA has been shown to be superior to AFM and BKT [2]. However, these findings have been disputed for BKT previously [29], [54]. Note that all models exhibit a much lower RMSE on the Algebra data set than on the other data sets. This better performance results from the bias in the data set towards correct observations (85% of correct observations): prediction accuracy is related to the bias of the data set, i.e., a uniform distribution of correct and wrong observations is most difficult to predict. The
significant reduction obtained in RMSE is very promising. The RMSE is a proper score [55] and has been widely used in educational data mining to measure student model performance, for example [7], [10], [11], [25]. The model fitting procedures we used optimize the log-likelihood of the data; the RMSE shows a high correlation to this measure [34]. Furthermore, the RMSE is also correlated to the ‘moment of learning’ (step in which a student learnt a skill) [35], [52].

We also computed the AUC of the different models. While the AUC was criticized as a performance measure lately [34], [35], it is a useful additional measure (for example in combination with the RMSE) to assess the models’ ability to discriminate failure from success. In a biased data set with a high base rate (i.e., a high frequency of correct outcomes), a model always outputting the base rate performs well regarding the RMSE, since the prediction of the base rate achieves perfect reliability. However, such a model exhibits a low AUC. The DBN models consistently show the highest AUC over all hierarchical data sets. For the Spelling data sets, DBN and PFA perform similarly. BKT shows a significantly lower AUC than the DBN models.

For a better understanding of model behavior, a more detailed quantification of model performance is useful. To answer our second research question Where do the differences in prediction accuracy between the different student models come from?, we therefore computed the Brier score of the models [34]. The BKT model consistently showed the lowest resolution (RES) of all models and was outperformed by the DBN models over all data sets. This result is not unexpected: The AUC is related to the Brier score [56] and the BKT models exhibited the lowest AUC across all data sets of our experiments. Furthermore, these results are in line with previous work which found that BKT models tend to have a low resolution (RES) [34]. We observe mixed results for the reliability: BKT exhibited a significantly better reliability than the DBN model on the Number Representation data set and a significantly worse reliability than DBN on the Algebra and Spelling data sets - reliabilities were similar on all other data sets. We therefore conclude that the differences in RMSE between BKT and DBN models are mainly due to the low resolution of BKT. The DBN models performed similarly or better in reliability and resolution than PFA (with the exception of the Spelling data set) and therefore the differences in RMSE between these two models probably arise from the combination of the two measures. The AFM consistently showed a bad reliability (worst model over all data sets), but performed very well regarding the resolution. The bad reliability of the AFM might stem from the fact, that AFM does not differentiate between correct and wrong observations. The superior resolution of the logistic regression models is probably due to their nature, i.e., they are fitting a learning curve over time. Our analyses demonstrate that a detailed quantification of model performance is helpful for understanding the properties of the different models.

Lately, the evaluation of student models based on small differences in performance metrics has been questioned [25]. However, previous work has demonstrated, that small differences in RMSE can have a large impact on over-practice [7], [40]. Furthermore, the improvement in performance metric has a significant impact on the expected number of practice opportunities needed [57] and the instructional policy of the system [1]. Other research discussed the use of a single number as a performance metric [34], [36], [37]. We therefore used the Teal framework [36] to answer our third research question: How do the different student modeling techniques influence instructional policies? Teal is threshold dependent and can be applied to all models used in our experiments.

The BKT models consistently showed a low effort over all data sets, scores were in range with the other models. Notice, however, that BKT demonstrated a very low variance in effort and scores both over time and over the different data sets. This low variance is caused by the low resolution (and AUC) of BKT. While the generally low efforts and scores might seem promising, the low variance limits instructional design, since the effort and score cannot be tuned by adapting the threshold. The DBN models tended to be more pessimistic than BKT, i.e., they generally required a higher effort to reach the same score. This is due to the fact that these models represent forgetting (\(p_f \geq 0\)), which leads to a slower increase of the predicted probability \(\hat{p}_i\). Nevertheless, the DBN models reached a similar or higher score than BKT with a reasonable effort \(E < 10\) (students with an effort \(E > 10\) can be considered to be wheel-spinning) [53]. The PFA models showed similar effort and scores as BKT for lower thresholds, but effort increased very fast with not much gain regarding the score. These results are in line with the observations made in previous work [1]. We conclude from our experiments that DBN models yield meaningful instructional policies relating to cognitive mastery.

From Fig. 7, it is very well visible, that Teal is threshold dependent. Different thresholds are optimal for the different models. This dependency on a threshold gives flexibility to instructional design, since the resulting policies can be tuned using the threshold. However, the optimal choice of threshold might be difficult, but could be done in a data-driven way by first determining a target score. Therefore, methods not requiring a threshold are an interesting alternative. The predictive similarity policy [1] for example stops when the predicted probabilities do not change anymore regardless of the student’s answer (correct or wrong).
from learning system with student models and instructional policies already in place, resulting in the introduction of a bias into the generated log data.

6 Conclusion

In this work, we showed that DBN models are well suited for representing student knowledge. We extensively evaluated our approach on five large data sets from different learning domains containing students of a wide age range. The results demonstrate that the incorporation of skill topologies yields to significant improvement in prediction accuracy over traditional student models. Furthermore, the more complex hierarchical DBN models also lead to meaningful and interpretable instructional policies. The use of the same parametrization and constraint sets for all experiments demonstrates that basic assumptions about learning hold across different learning domains. To conclude, our results show that modeling skill topologies is beneficial and easy to use, as even simple hierarchies and parametrization lead to significant improvements in prediction accuracy and instructional policies.

In the future, we would like to analyze the influence of the skill hierarchies, the different parameters and the reasons for the performance differences between the different approaches further. We furthermore plan to apply the individualization techniques used in BKT [7], [10], [11] to DBNs and compare the benefits of introducing skill hierarchies to the advantages of student-individualized modeling techniques. Moreover, we would like to explore the possibility of learning skill hierarchies from data.

Acknowledgments

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